Published in final edited form as:

Am J Obstet Gynecol. 2015 March ; 212(3): 355.e1–355.e7. doi:10.1016/j.ajog.2014.10.023.

Management of Fetal Malposition in the Second Stage of Labor: A Propensity Score Analysis

Abigail R. AIKEN, MB/BChir, MPH^a, Catherine E. AIKEN, MB/BChir, PhD^{b,*}, Medhat S. ALBERRY, MBBS, PhD^b, Jeremy C. BROCKELSBY, MBBS, PhD^b, and James G. SCOTT, PhD^c

^aPopulation Research Center and LBJ School of Public Affairs, University of Texas at Austin, Texas, USA

^bDepartment of Obstetrics and Gynaecology, University of Cambridge; NIHR Cambridge Comprehensive Biomedical Research Centre, CB2 2SW, UK

^cRed McCombs School of Business and Division of Statistics and Scientific Computation, University of Texas at Austin, Texas, USA

Abstract

Objective—To determine the factors associated with selection of rotational instrumental versus cesarean delivery to manage persistent fetal malposition, and to assess differences in adverse neonatal and maternal outcomes following delivery by rotational instruments versus cesarean section.

Study Design—We conducted a retrospective cohort study over a 5-year period in a tertiary UK obstetrics center. 868 women with vertex-presenting, single, live-born infants at term with persistent malposition in the second stage of labor were included. Propensity-score stratification was used to control for selection bias: the possibility that obstetricians may systematically select more difficult cases for cesarean section. Linear and logistic regression models were used to compare maternal and neonatal outcomes for delivery by rotational forceps or ventouse versus cesarean section, adjusting for propensity scores.

Results—Increased likelihood of rotational instrumental delivery was associated with lower maternal age (OR= 0.95 $p<0.01$), lower BMI (OR=0.94 $p<0.001$), lower birth-weight (OR=0.95 $p<0.01$), no evidence of fetal compromise at the time of delivery (OR=0.31 $p<0.001$), delivery during the daytime (OR= 1.45, $P<0.05$), and delivery by a more experienced obstetrician (OR=7.21 $p<0.001$). Following propensity score stratification, there was no difference by delivery

© 2014 Elsevier Inc. All rights reserved.

*Corresponding author: Telephone: +44(0)1223 336871 cema2@cam.ac.uk.

Study conducted at Addenbrooke's Hospital, Cambridge, UK

Disclosure Statement: The authors report no conflict of interest.

Reprints: Reprints will not be available.

Publisher's Disclaimer: This is a PDF file of an unedited manuscript that has been accepted for publication. As a service to our customers we are providing this early version of the manuscript. The manuscript will undergo copyediting, typesetting, and review of the resulting proof before it is published in its final citable form. Please note that during the production process errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.

method in the rates of delayed neonatal respiration, reported critical incidents, or low fetal arterial pH. Maternal blood loss was higher in the cesarean group ($295.8 \pm 48\text{ml}$ $p < 0.001$).

Conclusions—Rotational instrumental delivery is often regarded as unsafe. However, we find that neonatal outcomes are no worse once selection bias is accounted for, and that the likelihood of severe obstetric hemorrhage is reduced. More widespread training of obstetricians in rotational instrumental delivery should be considered, particularly in light of rising cesarean section rates.

Keywords

cesarean section; operative vaginal delivery; fetal malposition; delivery; intra-partum care

Introduction

Fetal head malposition in the second stage of labor is a significant risk factor for adverse maternal and neonatal outcomes, and is associated with high rates of both instrumental delivery and cesarean section [1]. While some women will spontaneously deliver a malpositioned fetus, most require obstetric intervention [2]. In cases of persistent malposition, the obstetrician must choose between a potentially difficult rotational instrumental delivery and a second-stage cesarean section.

Instrumental rotation of the fetal head has fallen out of favor in modern obstetric practice in much of the world, despite data showing low complication rates [3, 4]. It has recently been demonstrated that, while the majority of obstetricians considered rotation of the fetal head to be an acceptable intervention (97%), less than half (41%) had performed it within the previous year [5]. Second-stage cesarean section is an increasingly common alternative [6], but carries a significant burden of maternal morbidity [7, 8].

A small number of studies have compared the morbidity associated with different instruments used to effect rotational delivery, and have found low prevalence of adverse maternal and neonatal outcomes, as well as increased risk of some adverse events with emergency cesarean section [9–11]. However, any comparison of delivery outcomes by rotational instruments versus second-stage cesarean section must confront the possibility that obstetricians systematically select more difficult cases for cesarean section, thereby introducing a selection bias. This study has two main objectives: first, to illuminate the factors that make an attempt at rotational instrumental delivery more likely, by modeling the obstetrician's decision-making process; and second, to use propensity-score stratification to create comparable groups that allow differences in maternal and fetal outcomes by delivery type to be tested reliably.

Materials and Methods

A cohort of 25,886 women with vertex-presenting, single, live-born infants at term (37–42 completed weeks of gestation), aiming for vaginal delivery was identified over a 5-year period (Jan 2008–Oct 2013) in a single tertiary obstetrics center in the UK. A sub-cohort of 868 women was identified with a confirmed cephalic fetal malposition in the second stage of labor. Of these, 833 underwent either cesarean section ($n=534$) or successful instrumental

delivery (n=299), and 35 underwent failed instrumental delivery, followed by second-stage cesarean section.

Fetal malposition was defined as any cephalic position greater than 45 degrees from direct occipito-anterior [12], and was diagnosed by digital examination. The rate of malpositions delivered by each method did not vary significantly across the study years. Deliveries where the obstetrician performed manual rotation of the fetal head followed by direct instrumental delivery were not considered to be cases of persistent fetal malposition, and were not included in the analysis. The indications and procedures for instrumental delivery in our center are defined in the operative vaginal delivery guidance from the Royal College of Obstetricians and Gynaecologists (RCOG, UK) [12]. The classification of and indications for operative vaginal delivery are materially identical to the American College of Obstetricians and Gynecologists (ACOG) Practice Bulletin Number 17 on operative vaginal delivery [13].

Rotational instrumental delivery was carried out with either Kielland's forceps or ventouse. Ventouse devices available in the unit include posterior and rotational metal cups, silastic cups, and Kiwi Omnicups. Of the 334 successful instrumental deliveries, 62.0% (n=207) were conducted with Kielland's forceps and 38.0% (n=127) using ventouse.

Data regarding each woman's pregnancy, labor, and delivery were recorded by midwives shortly after birth, and were subsequently obtained from the hospital's Protos data-recording system. The database is regularly validated by a rolling program of audits where the original case notes are checked against the information recorded in the database. No patient-identifiable data were accessed during this research, which was performed as part of a provision-of-service study for the obstetrics center. Individual medical records were not accessed at any stage. Institutional Review Board approval was therefore not required.

Characteristics of the materno-fetal dyad were extracted from the database, including maternal age (at time of delivery), BMI (at first-trimester prenatal booking), parity (prior to delivery), ethnicity, and birth-weight to the nearest gram. Also recorded were the time between diagnosis of second stage and delivery (time fully dilated), and the instrument selected. Gestational age (measured by crown-rump length at first-trimester ultrasound) was recorded to the nearest week. Only cases where birth occurred within the interval 37–42 weeks' completed gestation were included. No adjustment was made for infants found to be small or large for gestational age. The indication for delivery was also classified into those where there was evidence of fetal compromise (including pathological fetal-heart tracing, abnormal fetal-blood sampling result, evidence of sepsis) and those where delivery was undertaken on other grounds (including failure to progress in second stage and maternal exhaustion). Deliveries were conducted under regional anesthesia (epidural or spinal), excepting a small number who required general anesthetic because of time constraints or failure of regional anesthesia during the procedure.

The level of experience of the obstetrician attempting delivery and the time at which the delivery took place were also recorded. Obstetricians were classified into three types using years of training as a proxy for experience. Type-1 and Type-2 obstetricians have 3–5 years

and 5–10 years of obstetric training, respectively. Type 3 obstetricians typically have >10 years of clinical obstetric experience. Our study was conducted in a unit where 2 obstetricians are available to perform instrumental deliveries or cesarean sections at any time. The first is typically a type-1 obstetrician, and is always supported by an immediately available doctor with >5 years obstetric training: a type-3 obstetrician during the day, or type-2 overnight. All obstetricians had training in at least one method of rotational instrumental delivery, in line with RCOG training requirements.

Delay in neonatal respiration was recorded where spontaneous respiration was not achieved within 1 minute of delivery. Umbilical cord blood was obtained immediately following delivery, and the arterial pH recorded. Correlation between arterial and venous pH was checked to confirm accuracy of the measurements. Arterial pH was categorized as ≥ 7.1 or < 7.1 [14]. A critical-incident form was generated at delivery in the case of any obstetric or neonatal emergency, including neonatal resuscitation, post-partum hemorrhage, shoulder dystocia, severe perineal trauma, maternal visceral injury, or any other event generating an obstetric emergency call. Maternal blood loss was measured by operating-room staff immediately after delivery, using suction blood collection and weighing of swabs and other pads. Blood loss was treated as a numerical variable to the nearest milliliter, and also categorized as minor ($< 1500\text{ml}$) or major ($\geq 1500\text{ml}$). Severe perineal trauma was defined as any disruption to the anal sphincter complex. Simple groupwise comparisons of these outcomes for women undergoing rotational instrumental versus cesarean delivery were carried out using either Student's t-test or the Mann-Whitney test for numerical data, and Pearson's chi-squared test for categorical data.

Any rigorous attempt to compare outcomes for the rotational-instrumental and cesarean-section groups is complicated by the fact that obstetricians may systematically select more difficult cases for cesarean section. This selection bias may involve physicians' own training and experience, their immediate concern for fetal well-being, and anticipated fetal weight. An extensive set of these assignment-related variables are available in our data set, allowing us to explicitly model the obstetrician's decision-making process. This allows us to use propensity-score stratification to adjust for factors that influence the decision to move towards cesarean section. Propensity-score stratification involves two stages. First, we build a statistical model for the treatment assignment (instrumental versus cesarean delivery), given a suitable set of predictors. The propensity score is the predicted probability of receiving the treatment derived from this first model. We then build a second set of models to estimate the effect of the treatment on each clinical outcome of interest, conditional on subjects' propensity scores. This approach generates a balanced cohort of subjects whose baseline characteristics will be statistically similar, regardless of treatment status. For the purpose of estimating treatment effects, it is typically more robust than standard regression modeling, and may be formally justified under the potential-outcomes framework for causal inference [15]. The effect of the covariates themselves on the clinical outcome is captured by the propensity score, and is never explicitly modeled.

For the purpose of estimating propensity scores, the 35 failed instrumental deliveries were included in the instrumental group, as the goal of this first-stage analysis was to model the physician's initial treatment decision. For the purpose of estimating treatment effects, we ran

two sets of second-stage analyses: one set with the 35 failed instrumental deliveries included, and one with them excluded.

Propensity scores were generated using a logistic-regression model predicting assignment to the instrumental-delivery group (the “treatment”). The regression model included seven covariates found to be significantly different between women undergoing rotational instrumental and cesarean section, and which were thought to be clinically relevant: maternal age, maternal BMI, parity, birth-weight, evidence of fetal compromise, time of delivery, and degree of experience of the delivering obstetrician. Although birth-weight is unknown before delivery, it has been included because it plausibly may be anticipated by the physician and it strongly predicts the decision to move to cesarean delivery.

The resulting propensity scores were then stratified by quintile [16], and the balance of covariates between cases of rotational instrumental delivery and cesarean section delivery checked within each quintile to verify that no significant differences remained. Adverse maternal and fetal outcomes were then modeled using linear and binary logistic regression, including the type of delivery and dummy variables for the propensity score quintiles as predictors. Findings were considered statistically significant at an alpha level of 0.05. Power calculations were performed by Monte Carlo simulation. All data analysis was conducted using the R statistical software package version 2.14.1.

Results

868 women with confirmed fetal malpositions in the second stage of labor were identified. 534 (61.5%) were delivered directly by second-stage cesarean section; and 334 (38.5%) had an attempted rotational instrumental delivery, 299 of which resulted in successful delivery, and 35 of which were converted to second-stage cesarean section. Characteristics of the maternal-fetal dyad were compared between the instrumental-delivery and cesarean-section groups (Table 1). Women in the cesarean-section group were more likely to be older ($p<0.01$), to have higher BMI ($p<0.001$), and to have babies with higher birth-weights ($p<0.01$). In terms of events surrounding delivery, women in the cesarean-section group were more likely to have had a delivery involving evidence of fetal compromise ($p<0.001$), to have been delivered during the night ($p<0.01$), and to have been delivered by a less experienced obstetrician ($p<0.001$).

Table 2 shows the results of the logistic regression predicting assignment to rotational instrumental delivery (the “treatment”). Lower birth-weight ($p<0.01$), lower maternal age ($p<0.01$), lower maternal BMI, ($p<0.001$), higher parity ($p<0.1$), absence of evidence of fetal compromise ($p<0.001$), delivery during the daytime ($p<0.05$), and increased experience of obstetrician ($p<0.001$) are all significant predictors of assignment to rotational instrumental delivery.

The balance of covariates between the two groups was then checked within each propensity-score quintile. There were no significant differences between the groups on any covariates in any of the quintiles. The results for the fourth quintile are shown as an example (Table 3); note that the propensity score quintiles are based on the combined groups. The

characteristics of the materno-fetal dyad are now much more similar across the two groups than they were before stratification (c.f. Table 1).

Unadjusted comparison of delivery outcomes (Table 4) showed that a higher percentage of deliveries in the cesarean-section group was associated with a critical incident at the time of delivery ($p<0.01$) and increased estimated blood loss (491.6ml v. 792.5ml, $p<0.001$). As expected, there were a number of infants with shoulder dystocia (2.7%) and severe maternal perineal trauma (3–4th degree tears; 6%) in the rotational-instrumental group. In the cesarean-section group, 9.4% required administration of general anesthesia.

Table 5 shows the associations between mode of delivery and adverse maternal and neonatal outcomes following propensity-score adjustment. As a robustness check, results are shown both with and without the failed instrumental deliveries included in the rotational-instrumental cohort, and are very similar in both cases. There were no differences between deliveries preformed by rotational instruments versus cesarean section in the time to neonatal respiration, reported critical incidents associated with the delivery, or likelihood of fetal umbilical arterial pH of ≤ 7.1 . The estimated blood loss was higher in the cesarean group by $295.8 \pm 48\text{ml}$ ($p<0.001$). In addition, the likelihood of sustaining $>1500\text{ml}$ estimated blood loss was lower in the instrumental group (OR 0.24, $p<0.01$). Power calculations showed that, at a Type-I error rate of 5%, we have at least 80% power to detect odds ratios outside the interval (0.54, 1.85) for binary outcomes, and to detect blood-loss effect sizes of at least 135 milliliters.

Comment

After propensity-score adjustment, instrumental delivery does not appear to be associated with worse maternal and neonatal outcomes. If anything, it offers a significantly lower risk than cesarean section of postpartum hemorrhage. We found no difference in delay to neonatal respiration following instrumental delivery, and no clinically significant difference in the risk of a low fetal arterial pH. We also demonstrate systematic differences between women who are assigned by obstetricians to rotational instrumental delivery versus second-stage cesarean section. These differences include lower birth-weight, lower maternal age, lower BMI and higher parity. Obstetricians are also more likely to undertake rotational instrumental delivery when they have more experience and when working during daylight hours (which may reflect the availability of immediate back-up from more experienced colleagues).

Our results are in general agreement with previous studies examining the maternal and neonatal risks of rotational instrumental delivery [9–11]. In addition to the outcomes reported here, these studies are reassuring regarding maternal outcomes, including duration of hospital stay [17] and obstetric anal sphincter injury [11]; and neonatal outcomes, including fetal injury [9, 10, 18]. Although we did not specifically examine it here, even in cases where instrumental deliveries are not successful, the outcomes of such attempts do not appear to be worse outcomes than proceeding directly to cesarean section [19].

The main strength of our study is its methodological robustness in addressing selection bias. Systematic differences between delivery groups are likely to affect any observational study,

complicating any attempt to compare maternal and neonatal outcomes using standard regression analysis. The use of propensity-score stratification, a technique that is becoming more widely used in obstetrics [20, 21], offers major advantages in this context. In particular, as long as the covariates can be shown to be properly balanced after stratification, the subsequent estimate of the treatment effect does not rely upon the precise mathematical relationship between the outcome and the covariates. This stands in strong contrast to standard regression analysis: when the covariates are heavily imbalanced between the groups, as they are in our data set, all treatment effects estimated by regression depend upon the specific form of the model, and are not robust to violations of standard assumptions, such as linearity, separability of covariate effects, and homoscedasticity [22, 23].

The main limitations of our study include the inability to distinguish between different fetal malpositions (occipito-posterior, occipito-transverse etc.), and the lack of sub-division of the cohort to distinguish between deliveries conducted using rotational forceps and ventouse. While these data are available, a further sub-cohort analysis has not been performed, as sample sizes would be insufficiently large to allow adequate propensity-score stratification between groups. Additionally, we did not have information about attempts at manual rotation, as these are not routinely recorded in our database. Another limitation is the lack of information about caput, molding, and station. We were thus unable to account for the influence of these subjective but important variables in the decision-making process. We were also unable to control for the presence of maternal diabetes. Our study documents adverse maternal and neonatal outcomes at birth; however, we lack the follow-up data to ensure that there is no excess of late adverse outcomes in either group. Existing data suggest that there is no evidence of increased adverse neurodevelopmental outcomes in neonates following instrumental delivery [24]. Additionally, data are collected shortly after delivery, and therefore do not include information on length of stay in hospital, either for the mother or the neonate.

The association we demonstrate between obstetrician experience and likelihood of proceeding to instrumental delivery likely reflects the difficulty of such deliveries and the experience required to undertake them with confidence. Indeed, others have noted the importance of operator experience in the safe use of Kielland's forceps [25] and that junior obstetricians are relatively more likely to use rotational ventouse rather than forceps [26]. We have shown elsewhere that obstetricians in their first 5 years of training are more likely to have unsuccessful instrumental deliveries than more experienced obstetricians [27]. Our findings imply that increased training and experience for trainee obstetricians is important, especially in light of rising cesarean section rates. Other studies have also recognized a need for improved training in instrumental delivery techniques [28]. While 'real-life' experience is desirable, simulator-based training has been developed and may help fulfill some learning needs [29].

The obstetrician's perception of the safety of the mother and fetus plays a major role in the decision to perform rotational instrumental delivery in the face of persistent malposition. Our analysis shows that higher maternal weight and age, the expectation that the fetus is large, and the presence of fetal distress all make the choice of cesarean section more likely. However, once we adjust for these factors, it does not appear that rotational instrumental

delivery is associated with a higher rate of adverse outcomes. A risk of shoulder dystocia is inherent in vaginal deliveries, as is the risk of severe perineal trauma. The obstetrician must carefully weigh these risks against the increased risk of maternal hemorrhage and of requiring general anesthesia with second-stage cesarean section. Rotational instrumental delivery, particularly by Kielland's forceps, has been all but abandoned in many obstetric practices. Yet the findings presented here suggest that there is room for further debate about the inclusion of rotational instruments in the clinical toolkit of modern obstetricians, especially in settings where cesarean section has become the default mode of delivery.

Acknowledgments

Sources of Financial Support: ARA is supported by an NICHD Predoctoral Fellowship under Grant No. 1 F31 HD079182-01, and by grant 5 R24 HD042849 awarded to the Population Research Center at The University of Texas at Austin. JGS is partially funded by a CAREER grant from the U.S. National Science Foundation (DMS-1255187).

References

1. Senecal J, Xiong X, Fraser WD. Effect of fetal position on second-stage duration and labor outcome. *Obstet Gynecol.* 2005 Apr; 105(4):763–72. [PubMed: 15802403]
2. Gardberg M, Leonova Y, Laakkonen E. Malpresentations--impact on mode of delivery. *Acta Obstet Gynecol Scand.* 2011 May; 90(5):540–2. [PubMed: 21501123]
3. Al-Suhel R, Gill S, Robson S, Shadbolt B. Kjelland's forceps in the new millennium. Maternal and neonatal outcomes of attempted rotational forceps delivery. *Aust N Z J Obstet Gynaecol.* 2009 Oct; 49(5):510–4. [PubMed: 19780735]
4. Healy DL, Quinn MA, Pepperell RJ. Rotational delivery of the fetus: Kielland's forceps and two other methods compared. *Br J Obstet Gynaecol.* 1982 Jul; 89(7):501–6. [PubMed: 7093162]
5. Phipps H, de Vries B, Lee PN, Hyett JA. Management of occiput posterior position in the second stage of labour: a survey of obstetric practice in Australia and New Zealand. *Aust N Z J Obstet Gynaecol.* 2012 Oct; 52(5):450–4. [PubMed: 22804812]
6. Unterscheider J, McMenamin M, Cullinane F. Rising rates of caesarean deliveries at full cervical dilatation: a concerning trend. *Eur J Obstet Gynecol Reprod Biol.* 2011 Aug; 157(2):141–4. [PubMed: 21470764]
7. McKelvey A, Ashe R, McKenna D, Roberts R. Caesarean section in the second stage of labour: a retrospective review of obstetric setting and morbidity. *J Obstet Gynaecol.* 2010 Apr; 30(3):264–7. [PubMed: 20373928]
8. Pergialiotis V, Vlachos DG, Rodolakis A, Haidopoulos D, Thomakos N, Vlachos GD. First versus second stage C/S maternal and neonatal morbidity: a systematic review and meta-analysis. *Eur J Obstet Gynecol Reprod Biol.* 2014 Jan 3.
9. Tempest N, Hart A, Walkinshaw S, Hapangama D. A re-evaluation of the role of rotational forceps: retrospective comparison of maternal and perinatal outcomes following different methods of birth for malposition in the second stage of labour. *BJOG.* 2013 Sep; 120(10):1277–84. [PubMed: 23906197]
10. Stock SJ, Josephs K, Farquharson S, Love C, Cooper SE, Kissack C, et al. Maternal and neonatal outcomes of successful Kielland's rotational forceps delivery. *Obstet Gynecol.* 2013 May; 121(5): 1032–9. [PubMed: 23635740]
11. Bahl R, Van de Venne M, Macleod M, Strachan B, Murphy DJ. Maternal and neonatal morbidity in relation to the instrument used for mid-cavity rotational operative vaginal delivery: a prospective cohort study. *BJOG.* 2013 Nov; 120(12):1526–32. [PubMed: 23924292]
12. Bahl, R.; Strachan, B.; Murphy, DJ. Greentop Guideline 62; Operative Vaginal Delivery. Royal College of Obstetricians and Gynaecologists; UK: 2011 Feb.

13. Operative vaginal delivery. Clinical management guidelines for obstetrician-gynecologists. American College of Obstetrics and Gynecology. *Int J Gynaecol Obstet*. 2001 Jul; 74(1):69–76. [PubMed: 11480404]
14. Yeh P, Emary K, Impey L. The relationship between umbilical cord arterial pH and serious adverse neonatal outcome: analysis of 51,519 consecutive validated samples. *BJOG*. 2012 Jun; 119(7): 824–31. [PubMed: 22571747]
15. Rosenbaum PR, Rubin DB. The central role of the propensity score in observational studies for causal effects. *Biometrika* [Original research]. 1983; 70(1):41–55.
16. D'Agostino RB Jr. Propensity score methods for bias reduction in the comparison of a treatment to a non-randomized control group. *Stat Med*. 1998 Oct 15; 17(19):2265–81. [PubMed: 9802183]
17. Schiff E, Friedman SA, Zolti M, Avraham A, Kayam Z, Mashiach S, et al. A matched controlled study of Kielland's forceps for transverse arrest of the fetal vertex. *J Obstet Gynaecol*. 2001 Nov; 21(6):576–9. [PubMed: 12521771]
18. Cardozo LD, Gibb DM, Studd JW, Cooper DJ. Should we abandon Kielland's forceps? *Br Med J (Clin Res Ed)*. 1983 Jul 30; 287(6388):315–7.
19. Revah A, Ezra Y, Farine D, Ritchie K. Failed trial of vacuum or forceps--maternal and fetal outcome. *Am J Obstet Gynecol*. 1997 Jan; 176(1 Pt 1):200–4. [PubMed: 9024114]
20. Gilbert SA, Grobman WA, Landon MB, Spong CY, Rouse DJ, Leveno KJ, et al. Elective repeat cesarean delivery compared with spontaneous trial of labor after a prior cesarean delivery: a propensity score analysis. *Am J Obstet Gynecol*. 2012 Apr; 206(4):311, e1–9. [PubMed: 22464069]
21. Roman L, Raffo JE, Zhu Q, Meghea CI. A statewide medicaid enhanced prenatal care program: impact on birth outcomes. *JAMA Pediatr*. 2014 Mar 1; 168(3):220–7. [PubMed: 24394980]
22. Hill, JL.; Reiter, JP.; Zanutto, EL. A comparisons of experimental and observational data analyses. In: Gelman, A.; Meng, XL., editors. *Applied Bayesian Modeling and Causal Inference From an Incomplete-Data Perspective*. Wiley; 2004. p. 44-6.
23. Rubin DB. Estimating causal effects from large data sets using propensity scores. *Annals of Internal Medicine*. 1997; 127:757–63. [PubMed: 9382394]
24. Bahl R, Patel RR, Swingler R, Ellis M, Murphy DJ. Neurodevelopmental outcome at 5 years after operative delivery in the second stage of labor: a cohort study. *Am J Obstet Gynecol*. 2007 Aug; 197(2):147, e1–6. [PubMed: 17689628]
25. Jain V, Guleria K, Gopalan S, Narang A. Mode of delivery in deep transverse arrest. *Int J Gynaecol Obstet*. 1993 Nov; 43(2):129–35. [PubMed: 7905427]
26. Tan KH, Sim R, Yam KL. Kielland's forceps delivery: is it a dying art? *Singapore Med J*. 1992 Aug; 33(4):380–2. [PubMed: 1411670]
27. Aiken CE, Aiken AR, Brocklesby JC, Scott JG. Factors Influencing the Likelihood of Instrumental Delivery Success. *Obstet Gynecol*. 2014; 123(4):796–803. [PubMed: 24785607]
28. Sau A, Sau M, Ahmed H, Brown R. Vacuum extraction: is there any need to improve the current training in the UK? *Acta Obstet Gynecol Scand*. 2004 May; 83(5):466–70. [PubMed: 15059160]
29. Moreau R, Pham MT, Brun X, Redarce T, Dupuis O. Simulation of an instrumental childbirth for the training of the forceps extraction: control algorithm and evaluation. *IEEE Trans Inf Technol Biomed*. 2011 May; 15(3):364–72. [PubMed: 21257384]

Table 1

Characteristics of the Maternal-Fetal Dyad for the Full Sample and Stratified by Decision to Rotate

Characteristic	All Patients (N=868)	Rotational Instrumental Delivery (n=334)	Second Stage Cesarean (n=534)
Maternal age (yrs)	30.5	29.8	31.0**
Maternal BMI (kg/m²)	25.5	24.4	26.2***
Gestation (wks)	39.9	39.9	39.8
Parity			
0	454 (52.3)	166 (49.7)	288 (53.9)
1+	414 (47.7)	168 (50.3)	246 (46.1)
Time fully dilated (min)	159.1	159.8	161.2
Epidural			
Yes	601 (69.2)	229 (68.6)	372 (69.7)
No	267 (30.8)	105 (31.4)	162 (30.3)
Obstetrician type			
1	405 (46.7)	104 (31.1)	301 (56.4)***
2	366 (42.2)	159 (47.6)	207 (38.8)
3	97 (11.2)	71 (21.3)	26 (4.9)
Birth weight (g)	3592	3532	3630**
Fetal Indication			
Yes	439 (50.6)	116 (34.7)	313 (58.6)***
No	429 (49.4)	218 (65.3)	221 (41.4)
Ethnicity			
White	784 (90.3)	303 (90.7)	481 (90.1)
Southeast Asian	53 (6.1)	21 (6.3)	32 (6.0)
Black	9 (1.0)	5 (1.5)	4 (0.7)
Chinese	8 (1.0)	1 (0.3)	7 (1.3)
Other	14 (1.6)	4 (1.2)	10 (1.9)
Time of Delivery			
Day	467 (53.8)	206 (61.7)	261 (48.9)
Night	401 (46.2)	128 (38.3)	273 (51.1)**

Figures are means or n, percentages in parentheses

**
p<0.01,***
p<0.001

Table 2

Factors Associated with Decision to Perform Rotational Instrumental Delivery (N=868)

Variable	Odds Ratio (95% CI)	
Maternal Age (yrs)	0.95**	(0.93–0.98)
Maternal BMI (kg/m ²)	0.94***	(0.91–0.97)
Parity		
0		ref
1+	1.36 [‡]	(1.00–1.85)
Obstetrician Type		
1		ref
2	2.49***	(1.79–3.48)
3	7.21***	(4.22–12.64)
Birth Weight (per 100g)	0.95**	(0.92–0.98)
Fetal Indication	0.31***	(0.23–0.43)
Time of Delivery		
Night		ref
Day	1.45*	(1.05–2.01)

Model coefficients are expressed as odds ratios and 95% confidence intervals (CI).

[‡] p<0.1,

* p<0.05,

** p<0.01,

*** p<0.001

Table 3

Balance of Matched Covariates after Propensity Score Stratification (Fourth Quintile)

Characteristic	Rotational Instrumental Delivery (n=81)	Second Stage Cesarean (n=92)	P
Maternal age (yrs)	29.4	29.8	0.60
Maternal BMI (kg/m²)	23.8	24.0	0.69
Parity			
0	42 (51.9)	46 (50.0)	0.81
1+	39 (48.1)	46 (50.0)	
Obstetrician type			
1	27 (33.3)	31 (33.7)	0.96
2	45 (55.6)	58 (63.0)	
3	9 (11.1)	3 (3.3)	
Birth weight (g)	3531.7	3592.4	0.10
Fetal Indication			
Yes	58 (71.6)	68 (73.9)	0.73
No	23 (28.4)	24 (26.1)	
Time of Delivery			
Day	48 (59.3)	52 (56.5)	0.72
Night	33 (40.7)	40 (43.5)	

Figures are means or n, percentages in parentheses

Table 4

Maternal and Neonatal Outcomes for the Full Sample and Stratified by Decision to Rotate

Characteristic	All Patients (N=868)	Rotational Instrumental Delivery (n=334)	Second Stage Cesarean (n=534)
Estimated Blood Loss (ml)	677.2	491.6	792.5***
Delayed Neonatal Respiration	86 (9.9)	27 (8.1)	59 (11.0)
Critical Incident Reported	89 (10.3)	23 (6.9)	66 (12.4)**
Umbilical Arterial pH <7.1	43 (5.0)	14 (4.2)	29 (5.4)
Shoulder Dystocia	9 (1.0)	9 (2.7)	-
Severe Perineal Trauma	20 (2.3)	20 (6.0)	-
General Anesthesia	50 (5.8)	-	50 (9.4)

Figures are means or n, percentages in parentheses

Failed Instrumentals included in rotational instrumental delivery group

**
p<0.01,***
p<0.001

Table 5

Associations Between Mode of Delivery and Adverse Maternal and Neonatal Outcomes for the Propensity Score Adjusted Sample

	Failed Instrumental Deliveries Not Included (N=833)		Failed Instrumental Deliveries Included (N=868)	
Outcome	OR (95% C.I.)	P	OR (95% C.I.)	P
Time to Neonatal Respiration				
Cesarean Section	ref		ref	
Rotational Instrumental	0.77 (0.43–1.31)	0.35	0.77 (0.44–1.29)	0.31
Incidence of Critical Incident				
Cesarean Section	ref		ref	
Rotational Instrumental	1.52 (0.77–3.09)	0.24	1.66 (0.86–3.31)	0.14
Fetal Umbilical Arterial pH<7.1				
Cesarean Section	ref		ref	
Rotational Instrumental	0.63 (0.26–1.40)	0.27	0.77 (0.36–1.60)	0.49
Estimated Blood Loss >1.5L				
Cesarean Section	ref		ref	
Rotational Instrumental	0.20 (0.10–0.38)	<0.01**	0.24 (0.13–0.43)	<0.01**
	Coefficient (S.E.)		Coefficient (S.E.)	
Estimated Blood Loss (ml)				
Cesarean Section	ref		ref	
Rotational Instrumental	–333.4 (50.09)	<0.001***	–295.8 (47.98)	<0.001***

**
p<0.01,

p<0.001