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Does Major Depressive Disorder Change with Age?

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Abstract

Objective—The authors used results from a twenty-year, high-intensity follow-up to measure the influence of aging, and of age at onset, on the long-term persistence of symptoms in major depressive disorder (MDD).

Method—Subjects who completed a 20-year series of semi-annual and then annual assessments with a stable diagnosis of MDD, or schizoaffective disorder other than mainly schizophrenic, ($n = 220$), were divided according to their ages at intake into youngest (18–29 years), middle (30–44 years), and oldest (45 years) groups. Depressive morbidity was quantified as the proportion of weeks spent in major depressive or schizoaffective episodes. General linear models (GLM) then tested for effects of time and time-by-group interactions on these measures. Regression analyses compared the influence of age of onset and of current age.

Results—Analyses revealed no significant time or group-by-time effects on the proportions of weeks in major depressive episodes in any of three age groups. Earlier ages of onset were associated with greater symptom persistence, particularly in the youngest group. The proportions of weeks ill showed intra-individual stability over time that was most evident in the oldest group.

Conclusion—These results indicate that the persistence of depressive symptoms in MDD does not change as individuals move from their third to their fifth decade, from their fourth to their sixth decade, or from their sixth to their eighth decade. An early age of onset, rather than youth per se, is associated with greater morbidity over two decades.

Keywords

major depression; age periods; age of onset; symptom persistence

Introduction

Investigations into the influence of age on the clinical expression of mood disorders typically rely on comparisons of ill individuals grouped by current age. This approach, however, confounds age of onset with current age and thus cannot be used to assess intra-

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individual changes over time. Age of onset in adolescence or early adulthood is strongly associated with high familial loadings for mood disorder (Mendlewicz and Baron, 1981; Weissman *et al.*, 1984; Weissman *et al.*, 1988; Kupfer *et al.*, 1989; Lyons *et al.*, 1998; Zisook *et al.*, 2007), while onset in late adulthood may be associated with atherosclerotic changes in the brain (Krishnan *et al.*, 1997; de Groot *et al.*, 2000), with early-onset Alzheimer's disease (Alexopoulos. *et al.*, 1993), and with other medical conditions that become more likely with age. Thus, the age period during which a mood disorder first appears may reflect fundamental differences in the pathophysiology of the illness and such differences may have sustained influences on the disorder's course. Once the disorder is manifest, though, its course may evolve as the individual ages, either through changes in life circumstances typical of passage from one age-period to another, or through age-related changes in the underlying pathophysiology of the disorder.

For instance, both a younger age-of-onset (Giles *et al.*, 1989) and a younger current age (Lindquist, 1945; Gonzales *et al.*, 1985; Coryell *et al.*, 1991) have been associated with higher recurrence rates in major depressive disorder (MDD), particularly with recurrences accompanied by stressors (Hammen *et al.*, 1992; Coryell *et al.*, 1994). Is the sort of MDD that begins in adolescence or early adulthood inherently more prone to recurrence than is MDD that begins later? Or do higher recurrence rates in youth result from physiological and/or environmental risk factors that are more likely during adolescence or early adulthood than later in life?

The possibility of age-related changes in morbidity has a more practical importance for patients who, understandably, would like to know the ways in which their illness is likely to change as they grow older. Unfortunately, the literature has very little to offer in answer to such questions. The requisite study must be sufficiently descriptive of symptom levels over time and have a duration that encompasses substantially different age periods. The numbers of subjects followed must also be large enough to allow for comparisons of separate age groupings. Mood disorders may well evolve differently in the transition from early adulthood to middle years than they do from the middle years to old age.

We are aware of only four efforts to prospectively follow individuals with mood disorder using multiple assessments over a period longer than ten years (Angst, 1998; Brodaty *et al.*, 2001; Angst *et al.*, 2003; Coryell *et al.*, 2003). None of these have examined long-term changes in symptom quality or intensity in separate age cohorts. The following report uses data from one of these four studies to determine whether patients with non-bipolar major depressive disorder show changes in depressive morbidity as they move from their mid-twenties to their mid-forties, from their late thirties to their late fifties, and from their late fifties to their mid-seventies. If an early onset indicates a depressive illness that is inherently more persistent or recurrent, then younger individuals should experience more morbidity than older individuals, both early in follow-up and as they age. If, instead, youth itself, rather than age at onset, predisposes to greater depressive morbidity, then the younger subjects should experience less morbidity as they age.

Methods

Subjects

The National Institute of Mental Health Collaborative Program of Psychobiology of Depression (CDS) recruited patients who met Research Diagnostic Criteria (RDC) (Spitzer *et al.*, 1978) for MDD, mania or schizoaffective disorder as they sought treatment at any of five academic centers from 1978 to 1981 (Coryell *et al.*, 2003). Because the study was originally designed to test genetic hypotheses, inclusion required participants to be white.

They were also required to be 18 years or older, knowledgeable of their biological parents, and English speaking. All participants provided informed consent.

The following analysis is limited to subjects who lacked any history of mania, hypomania or schizoaffective mania, or of schizoaffective depression, mainly schizophrenic subtype when they entered the study and who did not develop those syndromes during follow-up. Those who met RDC for the mainly affective or “other” subtype of schizoaffective disorder were retained for this analysis because the definition overlaps almost entirely with that of DSM-IV for MDD with mood-incongruent psychotic features. Subjects were further limited to those who completed twenty years of follow-up.

Procedures

Raters used the Schedule for Affective Disorders and Schizophrenia (SADS) (Endicott and Spitzer., 1978) to establish current and lifetime diagnoses at intake according to the RCD. SADS ratings integrated information from patient interview, informant interview when this was available, and medical record review.

Follow-up assessments occurred semiannually for the first five years following intake and annually thereafter. Raters used information from direct patient interview and from medical records to complete the Longitudinal Interval Follow-up Evaluation (LIFE) (Keller *et al.*, 1987) during the first two years, the LIFE II during years three through five, and the Streamlined Longitudinal Interval Continuation Evaluation (SLICE) in the sixth year and beyond. The latter two instruments differed from the LIFE in the omission of some psychosocial measures.

Each RDC syndrome that was active at intake, or that developed during follow-up, was thereafter assigned a weekly Psychiatric Symptom Rating (PSR). Interviewers identified change points in symptom intensity and quantified symptom levels for the interval using a six-point scale for MDD, mania, schizoaffective depression or schizoaffective mania. A score of `1' indicated no symptoms, `2' indicated the presence of no more than one or two symptoms to a mild degree, `5' indicated a full syndrome and `6' a relatively severe, full syndrome. Scores of `3' and `4' indicated the continued presence of an episode with less than the number of symptoms necessary for an initial diagnosis. The end of an episode required eight consecutive weeks of PSR ratings no greater than 2 and a new episode required that the subject again meet criteria for definite MDD or SA-D. Depressive morbidity was quantified as the percent of weeks in episodes of MDD or SA-D during prospective follow-up.

Data Analytic Procedures

Age at intake was used to assign patients into young (18–29 years), middle (age 30–44 years) and oldest (older than 44 years) groups. These age ranges were chosen beforehand to approximate contrasting age periods. Depressive morbidity over time was quantified as the percent of weeks during which an individual was in an episode of MDD or schizoaffective depression according to the convention that the PSR for either was greater than `2'. A one-way ANOVA was used to compare age-groups by morbidity levels. To determine whether depressive symptom morbidity changed differentially over time by age group, the twenty years of follow-up were grouped into five-year periods. The SAS General Linear Models (GLM) procedure tested whether age groups differed in any of the four follow-up periods. A repeated-measures analysis of variance then tested for time effects (differences across the five-year segments) and for time-by-group effects. Additional analyses tested for time-by-sex effects on both outcome measures. Because the distribution in percents of weeks ill varied across age-groupings and follow-up periods we also used the proportion of

individuals who were in episodes for greater than fifty percent of weeks in each five-year period to depict possible changes in morbidity levels over time.

To determine whether age of onset predicted depressive morbidity independent of current age we next introduced this variable into logistic regression models on the proportion of time in episodes over the twenty years of follow-up. Age of onset was then added to the previously described GLM models to determine whether it predicted changes over time in depressive morbidity in all subjects together or within any of the three age groups.

All correlations are presented as Spearman coefficients. A two-tailed alpha level of 0.05 was used for each of the tests reported below. No correction was made for multiple testing.

Results

Of 603 patients who met the above criteria for non-bipolar major affective disorder at study entry, 120 (19.9%) died within the next twenty years and 37 and 55 converted to Bipolar I and Bipolar II disorder, respectively. Of the remaining 391, 171 (43.7%) were lost to follow-up leaving 220 subjects who were followed at least 20 years.

Of patients who retained an RDC diagnosis of non-bipolar major depressive disorder at the end of twenty years of follow-up, 93 were less than 30 years old when they entered the study, 85 were 30 – 44 years of age, inclusive, and 42 were 45 years old or older (Table 1). These three groups did not differ by sex, the proportion that were inpatients, or by Global Assessment Scale (GAS) scores (Endicott *et al.*, 1976) for the week preceding intake.

The mean (SD) proportion of weeks in episodes of MDD or SA-D over the entire 20 years of follow-up were 27.5 (2.8), 32.8 (2.9) and 29.2 (3.5), respectively, for the young, middle and oldest groups ($F = .714$, $df = 2$, $p = .491$). With the three age groups combined, no consistent trend appeared to indicate increasing or decreasing morbidity across the four follow-up periods; mean (SD) values for percent of weeks ill were 31.8 (31.7), 26.9 (34.7), 29.3 (31.7) and 31.5 (39.7), respectively, for the first to last five year periods ($F = 2.4$, $df = 3$, 217 , $p = 0.0655$). As illustrated in Figure 1, there was likewise no consistent trend in morbidity level across the four follow-up periods in any of the three age groups. The GLM procedure showed no interaction of age group and follow-up period in percent time ill ($F = 0.34$, $df = 6$, 651 , $p = 0.8918$). Figure 2 shows that, likewise, no temporal trends emerged for increasing or decreasing morbidity when this was expressed as the proportion of patients in each year who experienced more than 50% of weeks ill.

Inclusion of sex in the above models revealed a time-by-sex interaction for percent of time ill ($F = 3.67$, $df = 3,214$, $p = 0.0131$). Further exploration revealed that this interaction was present in the youngest group ($F = 3.23$, $df = 3,273$, $p = 0.0229$) but was absent in the middle ($F = 1.16$, $df = 3,249$, $p = 0.3237$) and oldest groups ($F = 1.83$, $df = 3,120$, $p = 0.1447$). Within the youngest group both measures showed tendencies for depressive symptoms to increase in women and to decrease in men but neither of these were significant.

With age-groups combined age at intake was not predictive of subsequent symptom persistence ($F = 0.080$, $df = 1/218$, $p = 0.777$). With age of onset added to the regression model age at intake and age of onset had positive ($t = 2.514$, $p = 0.013$) and negative ($t = -3.087$, $p = 0.002$) relationships, respectively, with subsequent symptom persistence. Age of onset was negatively associated with percent time ill over the 20 years of follow-up in the young ($t = -2.48$, $p = 0.015$), but not in the middle ($t = -1.28$, $p = 0.203$) and oldest ($t = -1.46$, $p = 0.151$) groups though the direction of the relationship was consistent. Curve fitting applied to the plot of age-of-onset against time ill in the youngest group revealed a substantially better fit for a quadratic curve ($p = 0.006$) than for a linear one ($p = 0.015$) and

visual inspection indicated that morbidity increased rapidly with ages of onset of 18 and less. The 40 subjects who both entered the study before age 30, and who recalled an age of onset of 18 or earlier, showed no temporal trends over the four follow-up periods, though. Mean percent times depressed were 35.8 (33.4), 29.0 (36.0), 33.7 (40.5) and 33.8 (41.0) in the first through the last follow-up period, respectively.

Because the length of follow-up necessarily entailed sample attrition, the above analyses were repeated with the inclusion of 258 subjects who failed to complete 20 years of assessments, whether through death or refusal. Age groups did not differ significantly by mean proportions of time depressed in any of the four follow-up periods in this larger sample. Nor did the age groups differ by change in time depressed across the four follow-up periods ($F = 0.383$, $p = 0.683$). With the age groups combined, GLM analyses showed no significant within-subject change over the follow-up period ($F = 2.04$, $p = 0.106$).

Within each of the age-groups, correlations between five-year epochs in percent time ill decreased as intervals widened but remained highly significant (Table 2). Moreover, the young, middle and oldest groups had progressively greater inter-period correlations. Overall mean (SD) values for these were 0.44 (.15), 0.51 (0.13) and 0.66 (0.10), respectively (ANOVA $F = 4.4$, $df = 2/15$, $p = 0.032$) and the difference between the youngest and oldest groups was significant (Tukey's post-hoc $p = 0.029$).

Discussion

Time spent in major depressive episodes showed constancy over a 20-year period regardless of whether subjects were transitioning from their mid-twenties to their mid-forties, from their mid-thirties to their mid-fifties, or from their fifties to their seventies. Though this presentation obscures a considerable heterogeneity of courses among subjects, the overall absence of any pathoplastic effect of age on depressive morbidity is striking and suggests that, at least among patients who seek treatment at tertiary care centers, getting older neither worsens or improves symptom load in major depressive disorder.

A difference between men and women emerged for the youngest group such that men tended to improve with time and women tended to worsen. This was unexpected and, though a time-by-sex interaction existed, symptom persistence did not significantly change with age in either sex. Speculation on the meaning of this interaction thus should await replication in another sample.

As noted earlier, some reports have associated younger age with higher risks for relapse in MDD (Gonzales *et al.*, 1985; Giles *et al.*, 1989; Coryell *et al.*, 1991). Such findings imply that levels of depressive morbidity were higher over time in younger individuals if it is assumed that younger individuals did not have shorter episode durations. Whether that was so or not is unclear as these studies did not address the effects of age on episode duration, nor did they describe the proportions of time ill. The results of these earlier reports are therefore not inconsistent with those described here.

Symptom persistence showed significant within-individual stability over time in all three age groups but stability across time was greater in the oldest group than in the youngest. Thus, the tendency of an individual's depressive illness to persist or recur appears to declare itself early and to be an enduring quality. That this is least so in the youngest group suggests that environmental influences have greater importance in the course of illness in this age period.

One of the caveats necessary here concerns comparability to other samples. Approximately three-quarters of the individuals described were inpatients at tertiary academic centers when

they entered the study from 1979 through 1981. Regarding applicability to other inpatient samples it is reassuring that the overall mean percent time ill in our sample, 31.7%, closely resembles the percent time ill, 32.8%, that described the course of an inpatient sample recruited at a British academic center and followed for 8–11 years (Kennedy *et al.*, 2004). That sample also showed stability in group symptom-levels over time. Yet, the course of both samples may differ substantially from those of outpatients seen by private mental health providers, and perhaps more so from those treated by primary care providers, or those who do not seek treatment at all. There is evidence that the milder and more transient episodes that often result in no treatment are more common in younger individuals (Coryell *et al.*, 1995; Coryell *et al.*, 2002; Solomon *et al.*, 2005). If so, such episodes may become less frequent with age.

In addition, the recruiting centers for this study were geographically limited to the Midwest and the Northeast and, because the study was designed to explore the genetics of affective illness, participants were limited to Caucasians. The findings thus may not fully apply to a more geographically or ethnically diverse population. Finally, 20% of the subjects died within 20 years of study entry and another 44% were lost to follow-up. Non-random attrition could also have had an effect on our results.

Another caveat concerns the sensitivity to lower levels of depressive symptoms of the measures used in these analyses. We adopted conventional methods of tracking major depressive disorder over time in which eight weeks at PSR symptom levels of one or two indicated a recovery and a full major depressive syndrome was then necessary before PSR values higher than 2 could be recorded. Within these periods of “recovery” symptoms of intermittent depressive disorder or minor depression were doubtless active at times in a number of subjects. The existing literature, though, did not lead us to expect these milder depressive symptoms to differ by age in their persistence. Moreover, because these conventions have been extensively applied in previous studies, their use here facilitates the comparison of results across studies.

The control of treatment is not feasible in long-term studies but the medications taken by these subjects were carefully monitored and have been described in earlier reports (Keller *et al.*, 1986; Dawson *et al.*, 1998; Leon *et al.*, 2003; Solomon *et al.*, 2005). After the first year of follow-up many subjects took no antidepressants, or did so only intermittently. Moreover, the antidepressants being used were predominantly tricyclics (Coryell *et al.*, 1995; Coryell *et al.*, 2002). Selective serotonin re-uptake inhibitors (SSRIs) came into use only later in the observation period. While the results of some studies have indicated that the efficacy of tricyclic antidepressants differs from that of SSRIs in certain patient groups defined by age, sex or severity (Grigoriadis *et al.*, 2003; Joyce *et al.*, 2003; Parker *et al.*, 2003; Baca *et al.*, 2004), none of these findings have reached broad consensus. Thus, though our conclusions cannot be generalized with certainty to patients taking currently available antidepressants, there appears to be no particular reason that they should not apply.

In summary, data from this long-term, high-intensity follow-up indicate that age of onset, but not current age or aging itself, influences prospectively observed morbidity levels as measured by the proportions of weeks in MDD episodes. The persistence of symptoms in a given age period, though, predicts the persistence of symptoms many years later, particularly in older individuals.

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Figure 1.
Mean Proportions of Weeks in Depressive Episodes over Five-year Periods by Age Group

Figure 2.
Percent of Subjects in Episodes for Majority of Weeks

Table 1**Status at Intake by Age-Grouping**

	youngest	middle	oldest
n	93	85	42
number (%) female ¹	56 (60.2)	63 (74.1)	24 (57.1)
mean (SD) age at onset ²	19.8 (4.8)	27.9 (7.4)	41.5 (13.5)
number (%) inpatient	66 (71.0)	65 (76.5)	35 (83.3)
mean (SD) GAS score	43.2 (13.0)	44.4 (11.0)	43.2 (11.6)
mean (SD) age ³	24.5 (3.5)	35.6 (4.2)	55.4 (6.7)
mean (SD) age at end of follow-up ⁴	44.5 (3.4)	55.6 (4.2)	75.2 (6.8)

¹ $\chi^2 = 5.2$, df = 2, p = 0.075

²F = 106.6, df = 2, 217, p = 0.000

³F = 679.2, df = 2, 217, p = 0.000

⁴F = 679.2, df = 2, 217, p = 0.000

Table 2
Relationships between Subsequent Follow-up Periods in Percent Time Ill by Age Group

	youngest, years			middle, years			oldest, years		
	6-10	11-15	16-20	6-10	11-15	16-20	6-10	11-15	16-20
years 1-5	.597 [†]	.296 [*]	.289 [*]	.510 [†]	.364 [*]	.348 [†]	.663 [†]	.717 [†]	.569 [†]
years 6-10		.431 [†]	.384 [†]		.623 [†]	.525 [†]		.748 [†]	.499 [*]
years 11-15			.647 [†]			.676 [†]			.759 [†]

^{*} Spearman correlation, $p < .01$

[†] Spearman correlation, $p < .001$