Research Article

Variation in the Correlation of G + C Composition with Synonymous Codon Usage Bias among Bacteria

Haruo Suzuki, Rintaro Saito, and Masaru Tomita

Institute for Advanced Biosciences, Keio University, Yamagata 997-0017, Japan

Received 31 January 2007; Accepted 4 June 2007

Recommended by Teemu Roos

G + C composition at the third codon position (GC3) is widely reported to be correlated with synonymous codon usage bias. However, no quantitative attempt has been made to compare the extent of this correlation among different genomes. Here, we applied Shannon entropy from information theory to measure the degree of GC3 bias and that of synonymous codon usage bias of each gene. The strength of the correlation of GC3 with synonymous codon usage bias, quantified by a correlation coefficient, varied widely among bacterial genomes, ranging from -0.07 to 0.95 . Previous analyses suggesting that the relationship between GC3 and synonymous codon usage bias is independent of species are thus inconsistent with the more detailed analyses obtained here for individual species.

Copyright © 2007 Haruo Suzuki et al. This is an open access article distributed under the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

1. INTRODUCTION

Most amino acids can be encoded by more than one codon (i.e., a triplet of nucleotides); such codons are described as being synonymous and usually differ by one nucleotide in the third position. In many organisms, alternative synonymous codons are not used with equal frequency. Various factors have been proposed to contribute to synonymous codon usage bias, including G + C composition, replication strand bias, and translational selection [1]. Here, we focus on the contribution of G + C composition to synonymous codon usage bias.

G + C composition has been widely reported to be correlated with synonymous codon usage bias [2–11]. However, no quantitative attempt has been made to compare the extent of this correlation among different genomes. It would be useful to be able to quantify the strength of the correlation of G + C composition with synonymous codon usage bias in such a way that the estimates could be compared among genomes.

Different methods have been used to analyse the relationships between G + C composition and synonymous codon usage. Multivariate analysis methods, such as correspondence analysis [5–7] and principal component analysis [8], have been widely used to construct measures accounting for the largest fractions of the total variation in synony-

mous codon usage among genes. Carbone et al. [2, 3] used the codon adaptation index as a “universal” measure of dominating codon usage bias. The measures obtained by these methods can be interpreted as having different features (e.g., G + C composition bias, replication strand bias, and translationally selected codon bias), depending on the gene groups analyzed. Therefore, these methods would be useful for exploratory data analysis but not for the analysis of interest here. By contrast, measures such as the “effective number of codons” [10] and Shannon entropy from information theory [11] are well defined; these measures can be regarded as representing the degree of deviation from equal usage of synonymous codons, independently of the genes analyzed. Previous analyses of the relationships between G + C composition and synonymous codon usage bias using these measures have had two problems. First, these measures of synonymous codon usage bias have failed to take into account all three aspects of amino acid usage (i.e., the number of different amino acids, their relative frequency, and their codon degeneracy), and therefore are affected by amino acid usage bias, which may mask the effects directly linked to synonymous codon usage bias. Second, previous analyses have compared the “degree” of synonymous codon usage bias with G + C content [defined as $(G + C)/(A + T + G + C)$], and have therefore yielded a nonlinear U-shaped relationship (a gene with a very low or very high G + C content has a high degree of synonymous

codon usage bias) [9–11]; it is thus difficult to quantify the nonlinear relationship.

To overcome the first of these problems, we use the “weighted sum of relative entropy” (E_w) as a measure of synonymous codon usage bias [12]. This measure takes into account all three aspects of amino acid usage enumerated above, and indeed is little affected by amino acid usage biases. To overcome the second problem, we compare the degree of synonymous codon usage bias (E_w) with the degree of G + C content bias (entropy) instead of simply the G + C content; this step can provide a linear relationship. The strength of the linear relationship can be easily quantified by using a correlation coefficient.

The approach of quantifying the strength of the correlation of G + C composition with synonymous codon usage bias by using the entropy and correlation coefficient is applied to bacterial species for which whole genome sequences are available.

2. MATERIALS AND METHODS

2.1. Software

All analyses were conducted by using G-language genome analysis environment software [13], available at <http://www.g-language.org>. Graphs such as the histogram and scatter plot were generated in the R statistical computing environment [14], available at <http://www.r-project.org>.

2.2. Sequences

We tested data from 371 bacterial genomes (see Additional Table 1 for a comprehensive list (available online at <http://www2.bioinfo.ttc.keio.ac.jp/genome/haruo/BSB-ST1.pdf>)). Complete genomes in GenBank format [15] were downloaded from the NCBI repository site (<ftp://ftp.ncbi.nih.gov/genomes/Bacteria>). Protein coding sequences containing letters other than A, C, G, or T and those containing amino acids with residues less than their degree of codon degeneracy were discarded. From each coding sequence, start and stop codons were excluded.

2.3. Analyses

2.3.1. Measure of the degree of synonymous codon usage bias

The relative frequency of the j th synonymous codon for the i th amino acid (R_{ij}) is defined as the ratio of the number of occurrences of a codon to the sum of all synonymous codons:

$$R_{ij} = \frac{n_{ij}}{\sum_{j=1}^{k_i} n_{ij}}, \quad (1)$$

where n_{ij} is the number of occurrences of the j th codon for the i th amino acid, and k_i is the degree of codon degeneracy for the i th amino acid.

The degree of bias in synonymous codon usage of the i th amino acid (H_i) was quantified with a measure of uncertainty (entropy) in Shannon’s information theory [16]:

$$H_i = - \sum_{j=1}^{k_i} R_{ij} \log_2 R_{ij}, \quad (2)$$

H_i can take values from 0 (maximum bias where only one codon is used and all other synonyms are not present) to a maximum value $H_{i\max} = -k_i((1/k_i) \log_2(1/k_i)) = \log_2 k_i$ (no bias where alternative synonymous codons is used with equal frequency; that is, for every j , $R_{ij} = 1/k_i$).

The relative entropy of the i th amino acid (E_i) is defined as the ratio of the observed entropy to the maximum possible in the amino acid:

$$E_i = \frac{H_i}{H_{i\max}} = \frac{H_i}{\log_2 k_i}, \quad (3)$$

E_i ranges from 0 (maximum bias when $H_i = 0$) to 1 (no bias when $H_i = \log_2 k_i$).

To obtain an estimate of the overall bias in synonymous codon usage of a gene, we combined estimates of the bias from different amino acids, as follows. First, to take account of the difference in the degree of codon degeneracy (k_i) between different amino acids, we used the relative entropy (E_i) instead of the entropy (H_i) as an estimate of the bias of each amino acid. Second, to take account of the difference in relative frequency between different amino acids in the protein, we calculated the sum of the relative entropy of each amino acid weighted by its relative frequency in the protein. The measure of synonymous codon usage bias, designated as the “weighted sum of relative entropy” (E_w) [12], is given by

$$E_w = \sum_{i=1}^s w_i E_i, \quad (4)$$

where s is the number of different amino acid species in the protein and w_i is the relative frequency of the i th amino acid in the protein as a weighting factor. E_w ranges from 0 (maximum bias) to 1 (no bias).

2.3.2. Measure of the degree of G + C composition bias

The entropy was calculated to quantify the degree of bias in G + C composition at the first, second, and third codon positions of a gene (H_{GC1} , H_{GC2} , and H_{GC3} , resp.),

$$H_p = -p \log_2 p - (1 - p) \log_2 (1 - p), \quad (5)$$

where p is the G+C content (defined as (G+C)/(A+T+G+C)) at the first, second, or third codon positions in the nucleotide sequence (GC1, GC2, or GC3).

The entropy (H) for G + C composition (and for usage of two-fold degenerate codons; coding for asparagine, aspartic acid, cysteine, glutamic acid, glutamine, histidine, lysine, phenylalanine, or tyrosine) with values p and $1 - p$ is plotted in Figure 1 as a function of p .

FIGURE 1: Entropy (H) of G + C composition and usage of two fold degenerate codons with values p and $1 - p$.

2.3.3. Estimation of the correlation of G + C composition with synonymous codon usage bias

Spearman's rank correlation coefficient (r) was calculated to quantify the strength of the correlation between G + C composition bias (H_{GC1} , H_{GC2} , and H_{GC3}) and synonymous codon usage bias (E_w),

$$r = \frac{\sum_{g=1}^m (x_g - \bar{x})(y_g - \bar{y})}{\sqrt{\sum_{g=1}^m (x_g - \bar{x})^2 \sum_{g=1}^m (y_g - \bar{y})^2}}, \quad (6)$$

$$\bar{x} = \frac{1}{m} \sum_{g=1}^m x_g, \quad \bar{y} = \frac{1}{m} \sum_{g=1}^m y_g,$$

where x_g is the rank of the x -axis value (H_{GC1} , H_{GC2} , or H_{GC3}) for the g th gene, y_g is the rank of the y -axis value (E_w) for the g th gene, and m is the number of genes in the genome. The r value can vary from -1 (perfect negative correlation) through 0 (no correlation) to $+1$ (perfect positive correlation).

3. RESULTS

3.1. Correlation of G + C composition with synonymous codon usage bias (r value)

We investigated the correlation between the degree of G + C composition bias (H_{GC1} , H_{GC2} , and H_{GC3}) and that of synonymous codon usage bias (E_w) within each genome.

Figure 2 shows scatter plots of E_w plotted against H_{GC1} , H_{GC2} , and H_{GC3} with *Geobacter metallireducens* GS-15 genes and with *Saccharophagus degradans* 2–40 genes as examples and the Spearman's rank correlation coefficient (r) calculated from each plot. In *G. metallireducens*, the value of E_w was

much better correlated with H_{GC3} (Figure 2(c)) than with H_{GC1} (Figure 2(a)), or H_{GC2} (Figure 2(b)), indicating that GC3 contributed more to synonymous codon usage bias than GC1 and GC2. In *S. degradans*, the value of E_w was not correlated with H_{GC1} (Figure 2(d)), H_{GC2} (Figure 2(e)), or H_{GC3} (Figure 2(f)), indicating that neither GC1, nor GC2 nor GC3 contributed to synonymous codon usage bias.

To compare the contributions of GC1, GC2, and GC3 to synonymous codon usage bias, we produced pairwise scatter plots of the r values of H_{GC1} , H_{GC2} , and H_{GC3} with E_w for 371 genomes (Figure 3).

In the scatter plot of the r values of H_{GC3} (y -axis) plotted against those of H_{GC1} (x -axis) (Figure 3(a)), 362 points (97.6% of the total) are on the upper left of the line $y = x$, indicating that GC3 contributed more to synonymous codon usage bias than did GC1 in most of the genomes analyzed.

In the scatter plot of the r values of H_{GC3} (y -axis) plotted against those of H_{GC2} (x -axis) (Figure 3(b)), 367 points (98.9% of the total) are on the upper left of the line $y = x$, indicating that GC3 contributed more to synonymous codon usage bias than did GC2 in most genomes analyzed.

In the scatter plot of the r values of H_{GC1} (y -axis) plotted against those of H_{GC2} (x -axis) (Figure 3(c)), the scatter plot displays a diffuse distribution of points: 186 points (50.1% of the total) are on the upper left of the line $y = x$, indicating that the relative contributions of GC1 and GC2 to synonymous codon usage bias varied widely from genome to genome.

We constructed histograms showing the distribution of r values of H_{GC1} , H_{GC2} , and H_{GC3} with E_w for 371 bacterial genomes (Figure 4). The r values of H_{GC1} (Figure 4(a)) and H_{GC2} (Figure 4(b)) were distributed evenly between positive and negative values, whereas those of H_{GC3} (Figure 4(c)) were distributed towards positive values. The ranges [minimum, maximum] of the r values of H_{GC1} , H_{GC2} , and H_{GC3} were $[-0.51, 0.46]$, $[-0.28, 0.39]$, and $[-0.07, 0.95]$, respectively. The r values of H_{GC1} (Figure 4(a)) and H_{GC2} (Figure 4(b)) exhibited a monomodal distribution, whereas those of H_{GC3} (Figure 4(c)) exhibited a multimodal distribution.

3.2. Correlation of r value with genomic features

To investigate whether the correlation of GC3 with synonymous codon usage bias (the r value of H_{GC3} versus E_w) was related to species characteristics, we compared the r values with genomic features such as genomic G + C content and tRNA gene copy number. Among the 371 genomes analyzed here, genomic G + C content ranged from 23% to 73% and tRNA gene copy number varied from 28 to 145.

We constructed scatter plots of the r values of H_{GC3} with E_w plotted against genomic G + C content and tRNA gene copy number for 371 genomes (Figure 5). The relationship between the r value of H_{GC3} and the tRNA gene copy number was unclear (Figure 5(b)). In contrast, the r values of H_{GC3} tended to be high in G + C-poor or G + C-rich genomes, revealing a nonlinear relationship between the r value of H_{GC3} and genomic G + C content (Figure 5(a)). The highest r value

FIGURE 2: Scatter plots of E_w plotted against (a) H_{GC1} , (b) H_{GC2} , and (c) H_{GC3} for *Geobacter metallireducens* GS-15 genes and against (d) H_{GC1} , (e) H_{GC2} , and (f) H_{GC3} for *Saccharophagus degradans* 2–40 genes. The extent of the correlation between H_{GC1} , H_{GC2} , and H_{GC3} and E_w is represented by Spearman's rank correlation coefficient (r).

of H_{GC3} (0.95) was found in *G. metallireducens*, with a genomic G+C content of 60% (Figure 2(c)). The lowest r value of H_{GC3} (-0.07) was found in *S. degradans*, with a genomic G + C content of 46% (Figure 2(f)). The mean and standard

deviation of the r values of H_{GC3} for G + C-poor bacteria (with genomic G + C contents less than 40%) were 0.58 and 0.12, respectively. The corresponding values for G + C-rich bacteria (with genomic G + C contents greater than 60%)

FIGURE 3: Pairwise scatter plots of the r values of H_{GC1} , H_{GC2} and H_{GC3} with E_w for 371 bacterial genomes. Comparison of the correlation with E_w of (a) H_{GC3} and H_{GC1} , (b) H_{GC3} and H_{GC2} , and (c) H_{GC1} and H_{GC2} .

FIGURE 4: Histograms of the distribution of r values of (a) H_{GC1} , (b) H_{GC2} , and (c) H_{GC3} with E_w for 371 bacterial genomes.

were 0.86 and 0.04. Thus, the r values of H_{GC3} for G + C-poor bacteria tended to be lower than those for G + C-rich bacteria.

4. DISCUSSION

Other investigators have reported that G + C composition is correlated with synonymous codon usage bias in many organisms. However, no quantitative attempt has been made to compare the extent of this correlation among different genomes. Here, we quantified the strength of the correlation of G + C composition bias (H_{GC1} , H_{GC2} , and H_{GC3}) with synonymous codon usage bias (E_w) by using a correlation coefficient (r). This approach allowed us to quantitatively compare the strength of this correlation among different genomes.

FIGURE 5: Scatter plots of the r values of H_{GC3} with E_w plotted against (a) genomic G+C content and (b) tRNA gene number for 371 bacterial genomes.

In a previous analysis of the relationships between G + C composition and synonymous codon usage bias, Wan et al. [9] stated that “GC3 was the most important factor in codon bias among GC, GC1, GC2, and GC3.” This is quantitatively supported by the pairwise comparison of the r values of H_{GC1} , H_{GC2} , and H_{GC3} (Figure 3). However, the statement by Wan et al. that “GC3 is the key factor driving synonymous codon usage and that this mechanism is independent of species” differs from our conclusion that the strength of the correlation of GC3 with synonymous codon usage bias (the r value of H_{GC3}) varies widely among species (Figure 4(c)). This discordance appears to have arisen because Wan et al. combined the genes from different genomes into a single dataset for their analysis. This analysis of combined data from different genomes masks the presence of genomes in which the correlation of GC3 with synonymous codon usage bias is negligible (such as that of *S. degradans*; Figure 2(f)); the results are thus inconsistent with those of the more detailed analyses obtained here for individual genomes.

Three factors, G+C composition, replication strand bias, and translational selection, are well documented to shape synonymous codon usage bias [1].

First, in bacteria with extreme genomic G + C compositions (either G + C-rich or A + T-rich), synonymous codon usage could be dominated by strong mutational bias (toward G + C or A + T) [17, 18]. The data in Figure 5(a) indicate that, although genomic G + C content was nonlinearly correlated with the r value of H_{GC3} , there are some exceptions; for example, *Nanoarchaeum equitans* Kin4-M and *Mycoplasma genitalium* G37 had identical genomic G + C contents of 32% but very different r values of H_{GC3} (0.34 and 0.87, resp.), and *Thermococcus kodakarensis* KOD1 had a genomic G + C content of around 50% but a high r value of H_{GC3} (0.86). The existence of the outliers suggests that, although muta-

tional biases have a major influence on the correlation of GC3 with synonymous codon usage bias, other evolutionary factors may play a part. For example, horizontal gene transfer among bacteria with different genomic G + C content can contribute to intragenomic variation in G + C content [19, 20].

Second, the spirochaete *Borrelia burgdorferi* exhibits a strong base usage skew between leading and lagging strands of replication (generally inferred as reflecting strand-specific mutational bias): genes on the leading strand tend to preferentially use G- or T-ending codons [21]. The r values of H_{GC3} for genes on the leading and lagging strands are similar (0.65 and 0.63, resp.). This suggests that strand bias has little influence on the correlation of GC3 with synonymous codon usage bias in *B. burgdorferi*.

Third, in bacteria with more tRNA genes, synonymous codon usage could be subject to stronger translational selection [22]. Figure 5(b) shows that tRNA gene copy number was not correlated with the r value of H_{GC3} . This suggests that translational selection has little influence on the correlation of GC3 with synonymous codon usage bias. Sharp et al. [22] showed that the S value as a measure of translationally selected codon usage bias is highly correlated with tRNA gene copy number but is not correlated with genomic G + C content. Thus, the r value of H_{GC3} can be used as a measure complementary to the S value.

The most accepted hypothesis for the unequal usage of synonymous codons in bacterial genomes is that the unequal usage is the result of a very complex balance among different evolutionary forces (mutation and selection) [23]. The combined use of the r value and other methods (e.g., the S value) will improve our understanding of the relative contributions of different evolutionary forces to synonymous codon usage bias.

ABBREVIATIONS

A:	Adenine
T:	Thymine
G:	Guanine
C:	Cytosine
GC1:	G + C content at the first codon position
GC2:	G + C content at the second codon position
GC3:	G + C content at the third codon position
H_{GC1} :	Entropy of GC1
H_{GC2} :	Entropy of GC2
H_{GC3} :	Entropy of GC3
E_w :	Weighted sum of relative entropy
r :	Spearman's rank correlation coefficient

ACKNOWLEDGMENTS

The authors thank Dr Kazuharu Arakawa (Institute for Advanced Biosciences, Keio University) for his technical advice on the G-language genome analysis environment, and Kunihiko Baba (Faculty of Policy Management, Keio University) for his technical advice on the R statistical computing environment. This work was supported by the Ministry of Education, Culture, Sports, Science, and Technology of Japan Grant-in-Aid for the 21st Century Centre of Excellence (COE) Program entitled "Understanding and Control of Life via Systems Biology" (Keio University).

REFERENCES

- [1] M. D. Ermolaeva, "Synonymous codon usage in bacteria," *Current Issues in Molecular Biology*, vol. 3, no. 4, pp. 91–97, 2001.
- [2] A. Carbone, F. Kepes, and A. Zinovyev, "Codon bias signatures, organization of microorganisms in codon space, and lifestyle," *Molecular Biology and Evolution*, vol. 22, no. 3, pp. 547–561, 2005.
- [3] A. Carbone, A. Zinovyev, and F. Képès, "Codon adaptation index as a measure of dominating codon bias," *Bioinformatics*, vol. 19, no. 16, pp. 2005–2015, 2003.
- [4] R. D. Knight, S. J. Freeland, and L. F. Landweber, "A simple model based on mutation and selection explains trends in codon and amino-acid usage and GC composition within and across genomes," *Genome Biology*, vol. 2, no. 4, pp. research0010.1–research0010.13, 2001.
- [5] J. R. Lobry and A. Necşulea, "Synonymous codon usage and its potential link with optimal growth temperature in prokaryotes," *Gene*, vol. 385, pp. 128–136, 2006.
- [6] D. J. Lynn, G. A. C. Singer, and D. A. Hickey, "Synonymous codon usage is subject to selection in thermophilic bacteria," *Nucleic Acids Research*, vol. 30, no. 19, pp. 4272–4277, 2002.
- [7] G. A. C. Singer and D. A. Hickey, "Thermophilic prokaryotes have characteristic patterns of codon usage, amino acid composition and nucleotide content," *Gene*, vol. 317, no. 1–2, pp. 39–47, 2003.
- [8] H. Suzuki, R. Saito, and M. Tomita, "A problem in multivariate analysis of codon usage data and a possible solution," *FEBS Letters*, vol. 579, no. 28, pp. 6499–6504, 2005.
- [9] X.-F. Wan, D. Xu, A. Kleinhofs, and J. Zhou, "Quantitative relationship between synonymous codon usage bias and GC composition across unicellular genomes," *BMC Evolutionary Biology*, vol. 4, p. 19, 2004.
- [10] F. Wright, "The 'effective number of codons' used in a gene," *Gene*, vol. 87, no. 1, pp. 23–29, 1990.
- [11] B. Zeeberg, "Shannon information theoretic computation of synonymous codon usage biases in coding regions of human and mouse genomes," *Genome Research*, vol. 12, no. 6, pp. 944–955, 2002.
- [12] H. Suzuki, R. Saito, and M. Tomita, "The 'weighted sum of relative entropy': a new index for synonymous codon usage bias," *Gene*, vol. 335, no. 1–2, pp. 19–23, 2004.
- [13] K. Arakawa, K. Mori, K. Ikeda, T. Matsuzaki, Y. Kobayashi, and M. Tomita, "G-language genome analysis environment: a workbench for nucleotide sequence data mining," *Bioinformatics*, vol. 19, no. 2, pp. 305–306, 2003.
- [14] R Development Core Team, *R: a language and environment for statistical computing*, R Foundation for Statistical Computing, Vienna, Austria, 2006.
- [15] D. A. Benson, I. Karsch-Mizrachi, D. J. Lipman, J. Ostell, and D. L. Wheeler, "GenBank," *Nucleic Acids Research*, vol. 35, supplement 1, pp. D21–D25, 2007.
- [16] C. E. Shannon, "A mathematical theory of communication," *Bell System Technical Journal*, vol. 27, pp. 379–423, 1948.
- [17] A. Muto and S. Osawa, "The guanine and cytosine content of genomic DNA and bacterial evolution," *Proceedings of the National Academy of Sciences of the United States of America*, vol. 84, no. 1, pp. 166–169, 1987.
- [18] N. Sueoka, "On the genetic basis of variation and heterogeneity of DNA base composition," *Proceedings of the National Academy of Sciences of the United States of America*, vol. 48, no. 4, pp. 582–592, 1962.
- [19] S. Garcia-Vallve, A. Romeu, and J. Palau, "Horizontal gene transfer in bacterial and archaeal complete genomes," *Genome Research*, vol. 10, no. 11, pp. 1719–1725, 2000.
- [20] R. J. Grocock and P. M. Sharp, "Synonymous codon usage in *Pseudomonas aeruginosa* PA01," *Gene*, vol. 289, no. 1–2, pp. 131–139, 2002.
- [21] J. O. McInerney, "Replicational and transcriptional selection on codon usage in *Borrelia burgdorferi*," *Proceedings of the National Academy of Sciences of the United States of America*, vol. 95, no. 18, pp. 10698–10703, 1998.
- [22] P. M. Sharp, E. Bailes, R. J. Grocock, J. F. Peden, and R. E. Sockett, "Variation in the strength of selected codon usage bias among bacteria," *Nucleic Acids Research*, vol. 33, no. 4, pp. 1141–1153, 2005.
- [23] P. M. Sharp, M. Stenico, J. F. Peden, and A. T. Lloyd, "Codon usage: mutational bias, translational selection, or both?" *Biochemical Society Transactions*, vol. 21, no. 4, pp. 835–841, 1993.